

Sums of a Special Class of Generalized Stoykov Distributions¹

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Generalized Stoykov distributions are weighted versions of generalized Laplace distributions. By analogy with generalized Laplace distributions, two kinds of generalized Stoykov distributions are considered, one-sided and two-sided. The formulas for the sum of a special class of generalized Stoykov distributions are presented.

Key words: Laplace distribution, Stoykov distribution, generalized Laplace distribution, generalized Stoykov distribution.

1. Generalized Stoykov Distributions

Generalized Stoykov distribution can be considered as weighted generalized Laplace distribution by a polynomial weight function. We define generalized (double-side) Stoykov distribution as a distribution of random variable ξ with probability density function given by the equation

$$f_{\xi}(x) = C_D(\mu, \sigma, k, m, s, p) e^{-\left(\frac{|x-\mu|}{\sigma}\right)^k} \left(1 + \left|\frac{x-\mu}{s}\right|^p\right)^m, \\ \mu \in R, \sigma \in R^+, p \in N, m \in N, k \in N \quad (1)$$

We will denote this distribution as $\xi \in GSD(\mu, \sigma, k, m, s, p)$.

We define generalized (one-side) Stoykov distribution with normalized weight argument as a distribution of random variable ξ with probability density function given by the equation

$$f_{\xi}(x) = C_O(\mu, \sigma, k, m, s, p) e^{-\left(\frac{x-\mu}{\sigma}\right)^k} \left(1 + \left(\frac{x-\mu}{s}\right)^p\right)^m, \\ x \geq \mu, \mu \in R, \sigma \in R^+, p \in N, m \in N, k \in N \quad (2)$$

We will denote this distribution as $\xi \in GSO(\mu, \sigma, k, m, s, p)$.

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2. Some Special Cases of Generalized Stoynev Distributions

It is clear that generalized Laplace distributions are special case of generalized Stoynev distribution. We have $GSO(\mu, \sigma, k, 0, s, p) = GLO(\mu, \sigma, k)$ and $GSD(\mu, \sigma, k, 0, s, p) = GLO(\mu, \sigma, k)$. The special case $\xi \in GSO(0, \lambda^{-1}, 1, 1, 1, 1)$ is the distribution of Lindley. The special case $\xi \in GSO(0, \lambda^{-1}, 2, 1, 1, 1)$ is the distribution of Stoynev described in [1]. In this case $C_O(0, \lambda^{-1}, 2, 1, 1, 1) = \frac{\lambda^3}{\lambda^2 + 2\lambda + 2}$.

3. Sum of Two $GSO(0, \sigma, 1, m, 1, 1)$ Distributions

Here we will consider sum of two $GSO(0, \sigma, 1, m, 1, 1)$ distributions. Let $\xi \in GSO(0, \sigma_1, 1, m_1, 1, 1)$ and $\eta \in GSO(0, \sigma_2, 1, m_2, 1, 1)$. We will calculate the density of the sum $\phi = \xi + \eta$. We have

$$\begin{aligned} f_\xi(x) &= C_O(0, \sigma_1, 1, m_1, 1, 1) e^{\frac{-x}{\sigma_1}} (1+x)^{m_1} = \\ &= \frac{1}{\sigma_1 \sum_{k=1}^{m_1} k! \binom{m_1}{k} \sigma_1^k} e^{\frac{-x}{\sigma_1}} (1+x)^{m_1}, x > 0, \\ f_\xi(x) &= 0, x < 0 \end{aligned} \quad (3)$$

and

$$\begin{aligned} f_\eta(x) &= C_O(0, \sigma_2, 1, m_2, 1, 1) e^{\frac{-x}{\sigma_2}} (1+x)^{m_2} = \\ &= \frac{1}{\sigma_2 \sum_{k=1}^{m_2} k! \binom{m_2}{k} \sigma_2^k} e^{\frac{-x}{\sigma_2}} (1+x)^{m_2}, x > 0 \\ f_\eta(x) &= 0, x < 0. \end{aligned} \quad (4)$$

We have

$$\begin{aligned} f_\phi(u) &= \frac{1}{\sigma_1 \sum_{k=1}^{m_1} k! \binom{m_1}{k} \sigma_1^k} \frac{1}{\sigma_2 \sum_{k=1}^{m_2} k! \binom{m_2}{k} \sigma_2^k} * \\ &\quad * e^{\frac{-u}{\sigma_1}} \int_0^u e^{v(\frac{1}{\sigma_1} - \frac{1}{\sigma_2})} (1+u-v)^{m_1} (1+v)^{m_2} dv = \\ &= \frac{1}{\sigma_1 \sum_{k=1}^{m_1} k! \binom{m_1}{k} \sigma_1^k} \frac{1}{\sigma_2 \sum_{k=1}^{m_2} k! \binom{m_2}{k} \sigma_2^k} e^{\frac{-u}{\sigma_1}} I, u > 0, \end{aligned} \quad (5)$$

where

$$I = \int_0^u e^{v\left(\frac{1}{\sigma_1} - \frac{1}{\sigma_2}\right)} (1+u-v)^{m_1} (1+v)^{m_2} dv. \quad (6)$$

Now we will calculate the integral I . We will consider the case where $\sigma_1 = \sigma_2 = \sigma$. In this case,

$$\begin{aligned} I &= \int_0^u e^{v\left(\frac{1}{\sigma_1} - \frac{1}{\sigma_2}\right)} (1+u-v)^{m_1} (1+v)^{m_2} dv = \\ &= \int_0^u (1+u-v)^{m_1} (1+v)^{m_2} dv = \\ &= \int_0^u \sum_{k=1}^{m_1} \binom{m_1}{k} (1+u)^k (-1)^{m_1-k} v^{m_1-k} * \\ &\quad * \sum_{s=1}^{m_2} \binom{m_2}{s} v^s dv = \\ &= \sum_{k=1}^{m_1} \binom{m_1}{k} (1+u)^k (-1)^{m_1-k} * \\ &\quad * \sum_{s=1}^{m_2} \binom{m_2}{s} \int_0^u v^{m_1-k+s} dv = \\ &= \sum_{k=1}^{m_1} \binom{m_1}{k} (1+u)^k (-1)^{m_1-k} * \\ &\quad * \sum_{s=1}^{m_2} \binom{m_2}{s} \frac{v^{m_1-k+s+1}}{m_1-k+s+1} \Big|_0^u = \\ &= \sum_{k=1}^{m_1} \binom{m_1}{k} (1+u)^k (-1)^{m_1-k} * \\ &\quad * \sum_{s=1}^{m_2} \binom{m_2}{s} \frac{u^{m_1-k+s+1}}{m_1-k+s+1}. \end{aligned} \quad (7)$$

In the case when $\xi \in GSO(0, \sigma, 1, 2, 1, 1)$ and $\eta \in GSO(0, \sigma, 1, 2, 1, 1)$ we have sum of Stoynev distributions. In this case,

$$\begin{aligned} f_\xi(x) &= f_\eta(x) = C_O(0, \sigma, 1, 2, 1, 1) e^{\frac{-x}{\sigma}} (1+x)^2 = \\ &= \frac{1}{\sigma(1+\sigma+\sigma^2)} e^{\frac{-x}{\sigma}} (1+x)^2, x > 0 \end{aligned} \quad (8)$$

$$f_{\xi}(x) = f_{\eta}(x) = 0, x < 0.$$

Next,

$$\begin{aligned} f_{\phi}(u) &= \frac{1}{\sigma^2(1+\sigma+\sigma^2)^2} e^{\frac{-u}{\sigma}} * \\ &* \int_0^u (1+u-v)^2(1+v)^2 dv = \\ &= \frac{1}{\sigma^2(1+\sigma+\sigma^2)^2} e^{\frac{-u}{\sigma}} I, u > 0 \end{aligned} \quad (9)$$

where

$$\begin{aligned} I &= \int_0^u (1+u-v)^2(1+v)^2 dv = \\ &= \frac{30u + 60u^2 + 40u^3 + 10u^4 + u^5}{30}. \end{aligned} \quad (10)$$

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